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Accuracy (trueness and precision) of measurement methods and results —

Part 4:

Basic methods for the determination of the
trueness of a standard measurement method

*Exactitude (justesse et fidélité) des résultats et méthodes de mesure —
Partie 4: Méthodes de base pour la détermination de la justesse d'une
méthode de mesure normalisée*



Reference number
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Contents

Page

1	Scope	1
2	Normative references	1
3	Definitions	2
4	Determination of the bias of a standard measurement method by an interlaboratory experiment	2
4.1	The statistical model	2
4.2	Reference material requirements	2
4.3	Experimental design considerations when estimating the bias of a measurement method	3
4.4	Cross-references to ISO 5725-1 and ISO 5725-2	3
4.5	Required number of laboratories	3
4.6	Statistical evaluation	4
4.7	Interpretation of the results of the statistical evaluation	4
5	Determination of the laboratory bias of one laboratory using a standard measurement method	5
5.1	Carrying out the experiment	5
5.2	Cross-references to ISO 5725-1 and ISO 5725-2	6
5.3	Number of test results	6
5.4	Choice of reference materials	6
5.5	Statistical analysis	6
6	The report to, and the decisions to be taken by, the panel	7
6.1	Report by the statistical expert	7
6.2	Decisions by the panel	7
7	Utilization of trueness data	7

Annexes

A	Symbols and abbreviations used in ISO 5725	8
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B	Example of an accuracy experiment	10
B.1	Description of the experiment	10
B.2	Precision assessment	10
B.3	Trueness assessment	10
B.4	Further analysis	11
C	Derivation of equations	21
C.1	Equations (5) and (6) (see 4.5)	21
C.2	Equations (19) and (20) (see 5.3)	22
D	Bibliography	23

Foreword

ISO (the International Organization for Standardization) is a worldwide federation of national standards bodies (ISO member bodies). The work of preparing International Standards is normally carried out through ISO technical committees. Each member body interested in a subject for which a technical committee has been established has the right to be represented on that committee. International organizations, governmental and non-governmental, in liaison with ISO, also take part in the work. ISO collaborates closely with the International Electrotechnical Commission (IEC) on all matters of electrotechnical standardization.

Draft International Standards adopted by the technical committees are circulated to the member bodies for voting. Publication as an International Standard requires approval by at least 75 % of the member bodies casting a vote.

International Standard ISO 5725-4 was prepared by Technical Committee ISO/TC 69, *Applications of statistical methods*, Subcommittee SC 6, *Measurement methods and results*.

ISO 5725 consists of the following parts, under the general title *Accuracy (trueness and precision) of measurement methods and results*:

- *Part 1: General principles and definitions*
- *Part 2: Basic method for the determination of repeatability and reproducibility of a standard measurement method*
- *Part 3: Intermediate measures of the precision of a standard measurement method*
- *Part 4: Basic methods for the determination of the trueness of a standard measurement method*
- *Part 5: Alternative methods for the determination of the precision of a standard measurement method*
- *Part 6: Use in practice of accuracy values*

Parts 1 to 6 of ISO 5725 together cancel and replace ISO 5725:1986, which has been extended to cover trueness (in addition to precision) and intermediate precision conditions (in addition to repeatability and reproducibility conditions).

Annex A forms an integral part of this part of ISO 5725. Annexes B, C and D are for information only.

Introduction

0.1 ISO 5725 uses two terms “trueness” and “precision” to describe the accuracy of a measurement method. “Trueness” refers to the closeness of agreement between the arithmetic mean of a large number of test results and the true or accepted reference value. “Precision” refers to the closeness of agreement between test results.

0.2 General consideration of these quantities is given in ISO 5725-1 and so has not been repeated in this part of ISO 5725. ISO 5725-1 should be read in conjunction with all other parts of ISO 5725, including this part, because it gives the underlying definitions and general principles.

0.3 The “trueness” of a measurement method is of interest when it is possible to conceive of a true value for the property being measured. Although, for some measurement methods, the true value cannot be known exactly, it may be possible to have an accepted reference value for the property being measured; for example, if suitable reference materials are available, or if the accepted reference value can be established by reference to another measurement method or by preparation of a known sample. The trueness of the measurement method can be investigated by comparing the accepted reference value with the level of the results given by the measurement method. Trueness is normally expressed in terms of bias. Bias can arise, for example, in chemical analysis if the measurement method fails to extract all of an element, or if the presence of one element interferes with the determination of another.

0.4 Two measures of trueness may be of interest and both are considered in this part of ISO 5725.

- a) Bias of the measurement method: where there is a possibility that the measurement method may give rise to a bias, which persists wherever and whenever the measurement is done, then it is of interest to investigate the “bias of the measurement method” (as defined in ISO 5725-1). This requires an experiment involving many laboratories, very much as described in ISO 5725-2.
- b) Laboratory bias: measurements within a single laboratory can reveal the “laboratory bias” (as defined in ISO 5725-1). If it is proposed to undertake an experiment to estimate laboratory bias, then it should be realized that the estimate will be valid only at the time of the experiment. Further regular testing is required to show that the laboratory bias does not vary; the method described in ISO 5725-6 may be used for this.

Accuracy (trueness and precision) of measurement methods and results —

Part 4:

Basic methods for the determination of the trueness of a standard measurement method

1 Scope

1.1 This part of ISO 5725 provides basic methods for estimating the bias of a measurement method and the laboratory bias when a measurement method is applied.

1.2 It is concerned exclusively with measurement methods which yield measurements on a continuous scale and give a single value as the test result, although the single value may be the outcome of a calculation from a set of observations.

1.3 In order that the measurements are made in the same way, it is important that the measurement method has been standardized. All measurements are to be carried out according to that standard method.

1.4 Bias values give quantitative estimates of the ability of a measurement method to give the correct (true) result. When a value for the bias of a measurement method is quoted, together with a test result obtained by that method, there is an implication that the same characteristic is being measured in exactly the same way.

1.5 This part of ISO 5725 can be applied only if the accepted reference value can be established as a conventional true value, for example by measurement standards or suitable reference materials or by refer-

ring to a reference measurement method or by preparation of a known sample.

Reference materials could be either

- a) certified reference materials;
- b) materials manufactured for the purpose of the experiment with known properties; or
- c) materials whose properties have been established by measurements using an alternative measurement method whose bias is known to be negligible.

1.6 This part of ISO 5725 considers only those cases where it is sufficient to estimate bias on one level at a time. It is not applicable if the bias in the measurement of one property is affected by the level of a second property (i.e. it does not consider interferences). Comparison of the trueness of two measurement methods is considered in ISO 5725-6.

NOTE 1 In this part of ISO 5725, bias is considered only at one level at a time. Therefore the index j for the level has been omitted throughout.

2 Normative references

The following standards contain provisions which, through reference in this text, constitute provisions of this part of ISO 5725. At the time of publication, the

editions indicated were valid. All standards are subject to revision, and parties to agreements based on this part of ISO 5725 are encouraged to investigate the possibility of applying the most recent editions of the standards indicated below. Members of IEC and ISO maintain registers of currently valid International Standards.

ISO 3534-1:1993, *Statistics — Vocabulary and symbols — Part 1: Probability and general statistical terms*.

ISO 5725-1:1994, *Accuracy (trueness and precision) of measurement methods and results — Part 1: General principles and definitions*.

ISO 5725-2:1994, *Accuracy (trueness and precision) of measurement methods and results — Part 2: Basic method for the determination of repeatability and reproducibility of a standard measurement method*.

3 Definitions

For the purposes of this part ISO 5725, the definitions given in ISO 3534-1 and in ISO 5725-1 apply.

The symbols used in ISO 5725 are given in annex A.

4 Determination of the bias of a standard measurement method by an interlaboratory experiment

4.1 The statistical model

In the basic model described in subclause 5.1 of ISO 5725-1:1994, the general mean m may be replaced by

$$m = \mu + \delta \quad \dots (1)$$

where

μ is the accepted reference value of the property being measured;

δ is the bias of the measurement method.

The model becomes

$$y = \mu + \delta + B + e \quad \dots (2)$$

Equation (2) is used when δ is of interest. Here B is the laboratory component of bias, i.e. the component in a test result representing the between-laboratory variation.

The laboratory bias, Δ , is given by

$$\Delta = \delta + B \quad \dots (3)$$

so the model may be written

$$y = \mu + \Delta + e \quad \dots (4)$$

Equation (4) is used when Δ is of interest.

4.2 Reference material requirements

If reference materials are used, the requirements given in 4.2.1 and 4.2.2 shall be satisfied. Reference materials shall be homogeneous.

4.2.1 Choice of reference materials

4.2.1.1 The reference material shall have known properties at the level appropriate to the level at which the standard measurement method is intended to be applied, e.g. concentration. In some cases it will be important to include, in the assessment experiment, a series of reference materials, each corresponding to a different level of the property, as the bias of the standard measurement method may be different at different levels. The reference material should have a matrix as close as possible to the matrix of the material to be subjected to the standard measurement method, e.g. carbon in coal or carbon in steel.

4.2.1.2 The quantity of the reference material shall be sufficient for the entire experimental programme, including some in reserve if this is considered necessary.

4.2.1.3 Wherever possible, the reference material should have stable properties throughout the experiment. There are three cases, as follows.

- The properties are stable: no precautions are necessary.
- The certified value of the property may be influenced by storage conditions: the container should be stored, both before and after its opening, in the way described on the certificate.
- The properties change at a known rate: there is a certificate supplied with the reference value to define the properties at specific times.

4.2.1.4 The possible difference between the certified value and the true value expressed by the uncertainty of the reference material (see ISO Guide 35) is not taken into account in the methods given here.

4.2.2 Check and distribution of the reference material

Where sub-division of the unit of the reference material occurs prior to distribution, it shall be performed with care to avoid the introduction of any additional error. Relevant International Standards on sample division should be consulted. The units should be selected on a random basis for distribution. If the measurement process is non-destructive, it is possible to give all the laboratories in the interlaboratory experiment the same unit of reference material, but this will extend the time-frame of the experiment.

4.3 Experimental design considerations when estimating the bias of a measurement method

4.3.1 The objective of the experiment is to estimate the magnitude of the bias of the measurement method and to determine if it is statistically significant. If the bias is found to be statistically insignificant, then the objective is to determine the magnitude of the maximum bias that would, with a certain probability, remain undetected by the results of the experiment.

4.3.2 The layout of the experiment is almost the same as that for a precision experiment, as described in subclause 4.1 of ISO 5725-2:1994. The differences are

- there is an additional requirement to use an accepted reference value, and
- the number of participating laboratories and the number of test results shall also satisfy the requirements given in 4.5.

4.4 Cross-references to ISO 5725-1 and ISO 5725-2

Clause 6 of ISO 5725-1:1994 and clauses 5 and 6 of ISO 5725-2:1994 apply. When reading parts 1 and 2 in this context, "trueness" should be inserted in place of "precision" or "repeatability and reproducibility" as appropriate.

4.5 Required number of laboratories

The number of laboratories and the number of test results required at each level are interdependent. The

number of laboratories to be used is discussed in subclause 6.3 of ISO 5725-1:1994. A guide to deciding how many is given below.

In order for the results of an experiment to be able to detect with a high probability (see annex C) a predetermined magnitude of bias, the minimum number of laboratories, p , and test results, n , shall satisfy the following equation:

$$A\sigma_R \leq \frac{\delta_m}{1,84} \quad \dots (5)$$

where

δ_m is the predetermined magnitude of bias that the experimenter wishes to detect from the results of the experiment;

σ_R is the reproducibility standard deviation of the measurement method.

A is a function of p and n and is given by

$$A = 1,96 \sqrt{\frac{n(\gamma^2 - 1) + 1}{\gamma^2 pn}} \quad \dots (6)$$

where

$$\gamma = \sigma_R / \sigma_r \quad \dots (7)$$

Values of A are given in table 1.

Ideally, the choice of the combination of the number of laboratories and the number of replicate test results per laboratory should satisfy the requirement described by equation (5), with the δ_m value predetermined by the experimenter. However, for practical reasons, the choice of the number of laboratories is usually a compromise between the availability of resources and the desire to reduce the value of δ_m to a satisfactory level. If the reproducibility of the measurement method is poor, then it will not be practical to achieve a high degree of certainty in the estimate of the bias. When σ_R is larger than σ_r (i.e. γ is larger than 1) as is often the case, little is to be gained by obtaining more than $n = 2$ test results per laboratory per level.

Table 1 — Values showing the uncertainty in the estimate of the bias of the measurement method

<i>p</i>	<i>γ</i> = 1			<i>γ</i> = 2			<i>γ</i> = 5		
	<i>n</i> = 2	<i>n</i> = 3	<i>n</i> = 4	<i>n</i> = 2	<i>n</i> = 3	<i>n</i> = 4	<i>n</i> = 2	<i>n</i> = 3	<i>n</i> = 4
5	0,62	0,51	0,44	0,82	0,80	0,79	0,87	0,86	0,86
10	0,44	0,36	0,31	0,58	0,57	0,56	0,61	0,61	0,61
15	0,36	0,29	0,25	0,47	0,46	0,46	0,50	0,50	0,50
20	0,31	0,25	0,22	0,41	0,40	0,40	0,43	0,43	0,43
25	0,28	0,23	0,20	0,37	0,36	0,35	0,39	0,39	0,39
30	0,25	0,21	0,18	0,33	0,33	0,32	0,35	0,35	0,35
35	0,23	0,19	0,17	0,31	0,30	0,30	0,33	0,33	0,33
40	0,22	0,18	0,15	0,29	0,28	0,28	0,31	0,31	0,31

4.6 Statistical evaluation

The test results shall be treated as described in ISO 5725-2. In particular, if outlying values are detected, all necessary steps shall be taken to investigate the reasons why they have been obtained, including re-appraisal of the suitability of the accepted reference value.

4.7 Interpretation of the results of the statistical evaluation

4.7.1 Check of precision

The precision of the measurement method is expressed in terms of s_r (estimate of the repeatability standard deviation) and s_R (estimate of the reproducibility standard deviation). Equations (8) to (10) assume an equal number (n) of test results in each laboratory. If this is not true, the respective equations given in ISO 5725-2 should be used to calculate s_r and s_R .

4.7.1.1 The estimate s_r^2 of the repeatability variance for p participating laboratories is calculated as

$$s_r^2 = \frac{1}{p} \sum_{i=1}^p s_i^2 \quad \dots (8)$$

$$s_i^2 = \frac{1}{n-1} \sum_{k=1}^n (y_{ik} - \bar{y}_i)^2 \quad \dots (9)$$

$$\bar{y}_i = \frac{1}{n} \sum_{k=1}^n y_{ik} \quad \dots (10)$$

where s_i^2 and \bar{y}_i are respectively the variance and the average of n test results y_{ik} obtained in laboratory i .

Cochran's test, as described in ISO 5725-2, shall be applied to the variances s_i^2 to verify that no significant

differences exist between the within-laboratory variances. Mandel's h and k plots, as described in ISO 5725-2, should also be drawn for a more thorough investigation of potential outliers.

If the repeatability standard deviation of the standard measurement method has not been previously determined in accordance with ISO 5725-2, s_r will be considered to be the best estimate of it. If the repeatability standard deviation of the standard test method, σ_r , has been determined in accordance with ISO 5725-2, s_r^2 can be assessed by computing the ratio

$$C = s_r^2 / \sigma_r^2 \quad \dots (11)$$

The test statistic C is compared with the critical value

$$C_{\text{crit}} = \chi_{(1-\alpha)}^2(v)/v$$

where $\chi_{(1-\alpha)}^2(v)$ is the $(1-\alpha)$ -quantile of the χ^2 distribution with $v [= p(n-1)]$ degrees of freedom. Unless otherwise stated, α is assumed to be 0,05.

- If $C \leq C_{\text{crit}}$, s_r^2 is not significantly larger than σ_r^2 .
- If $C > C_{\text{crit}}$, s_r^2 is significantly larger than σ_r^2 .

In the former case, the repeatability standard deviation, σ_r , will be used for the assessment of the bias of the measurement method. In the latter case, it is necessary to investigate the causes of the discrepancy and possibly to repeat the experiment prior to proceeding further.

4.7.1.2 The estimate, s_R^2 , of the reproducibility variance for the p participating laboratories, is calculated as

$$s_R^2 = \frac{1}{p-1} \sum_{i=1}^p (\bar{y}_i - \bar{\bar{y}})^2 + \left(1 - \frac{1}{n}\right) s_r^2 \quad \dots (12)$$

with

$$\bar{y} = \frac{1}{p} \sum_{i=1}^p \bar{y}_i \quad \dots (13)$$

If the reproducibility standard deviation of the standard measurement method has not previously been determined in accordance with ISO 5725-2, s_R will be considered the best estimate of it. If the reproducibility standard deviation, σ_R , and the repeatability standard deviation, σ_r , of the standard measurement method have been determined in accordance with ISO 5725-2, s_R can be assessed indirectly by computing the ratio

$$C' = \frac{s_R^2 - (1 - 1/n)s_r^2}{\sigma_R^2 - (1 - 1/n)\sigma_r^2} \quad \dots (14)$$

The test statistic C' is compared with the critical value

$$C'_{\text{crit}} = \chi^2_{(1-\alpha)}(v)/v$$

where $\chi^2_{(1-\alpha)}(v)$ is the $(1-\alpha)$ -quantile of the χ^2 distribution with v ($= p - 1$) degrees of freedom. Unless otherwise stated, α is assumed to be 0,05.

- If $C' \leq C'_{\text{crit}}$: $s_R^2 - (1 - 1/n)s_r^2$ is not significantly larger than $\sigma_R^2 - (1 - 1/n)\sigma_r^2$.
- If $C' > C'_{\text{crit}}$: $s_R^2 - (1 - 1/n)s_r^2$ is significantly larger than $\sigma_R^2 - (1 - 1/n)\sigma_r^2$.

In the former case, the repeatability standard deviation, σ_r , and the reproducibility standard deviation, σ_R , will be used for the assessment of the trueness of the measurement method. In the latter case, a careful examination of the working conditions of each laboratory shall be carried out before the assessment of the bias of the standard measurement method is undertaken. It may appear that some laboratories did not use the required equipment or did not work according to the specified conditions. In chemical analysis, problems may arise from, for example, insufficient control of temperature, moisture, presence of contaminants, etc. As a result the experiment may have to be repeated to yield the expected precision values.

4.7.2 Estimation of the bias of the standard measurement method

The estimate of the bias from the assessing laboratories is given by

$$\hat{\delta} = \bar{y} - \mu \quad \dots (15)$$

where $\hat{\delta}$ may be positive or negative.

If the absolute value of the estimated bias is smaller than or equal to half the width of the uncertainty interval, as defined in ISO Guide 35, there is no evidence of a bias.

The variation of the estimate of the bias of the measurement method is due to the variation in the results of the measurement process and is expressed by its standard deviation computed as

$$\sigma_{\hat{\delta}} = \sqrt{\frac{\sigma_R^2 - (1 - 1/n)\sigma_r^2}{p}} \quad \dots (16)$$

in the case of known precision values, or

$$s_{\hat{\delta}} = \sqrt{\frac{s_R^2 - (1 - 1/n)s_r^2}{p}} \quad \dots (17)$$

in the case of unknown precision values.

An approximate 95 % confidence interval for the bias of the measurement method can be computed as

$$\hat{\delta} - A\sigma_R \leq \delta \leq \hat{\delta} + A\sigma_R \quad \dots (18)$$

where A is as given in equation (6). If σ_R is unknown, its estimate s_R has to be used instead, and A has to be computed with $\gamma = s_R/s_r$.

If this confidence interval covers the value zero, the bias of the measurement method is insignificant at the significance level $\alpha = 5$ %; otherwise it is significant.

5 Determination of the laboratory bias of one laboratory using a standard measurement method

As described below, experiments in one laboratory are used to estimate laboratory bias, provided that an interlaboratory precision experiment, in accordance with ISO 5725-2, has established the repeatability standard deviation of the method.

5.1 Carrying out the experiment

The experiment shall conform strictly to the standard method and measurements shall be carried out under repeatability conditions. Prior to conducting the assessment of trueness, a check of the precision of the standard measurement method as applied by the laboratory shall be performed. This implies comparison between the within-laboratory standard deviation and the stated repeatability standard deviation of the standard measurement method.

The layout of the experiment consists of the measurements required of one laboratory in a precision experiment as described in ISO 5725-2. Apart from the restriction to a single laboratory, the only substantial difference is the additional requirement to use an accepted reference value.

When attempting to measure the bias of a laboratory, it may not be worth putting a great deal of effort into such an experiment: the effort could perhaps be better expended by making checks at intervals as described in ISO 5725-6. If the repeatability of the measurement method is poor, then it will not be practical to achieve a high degree of certainty in the estimate of the bias of the laboratory.

5.2 Cross-references to ISO 5725-1 and ISO 5725-2

When reading ISO 5725-1 and ISO 5725-2 in this context, "trueness" should be inserted in place of "precision" or "repeatability and reproducibility" as appropriate. In ISO 5725-2, the number of laboratories will be $p = 1$, and it may be convenient for one person to combine the roles of "executive" and "supervisor".

5.3 Number of test results

The uncertainty in the estimate of the laboratory bias depends on the repeatability of the measurement method and on the number of test results obtained.

In order for the results of an experiment to be able to detect with a high probability (see annex C) a predetermined magnitude of bias, the number of test results, n , shall satisfy the following equation:

$$A_W \sigma_r \leq \frac{\Delta_m}{1,84} \quad \dots (19)$$

where

Δ_m is the predetermined magnitude of laboratory bias that the experimenter wishes to detect from the results of the experiment;

σ_r is the repeatability standard deviation of the measurement method and

$$A_W = \frac{1,96}{\sqrt{n}} \quad \dots (20)$$

5.4 Choice of reference materials

If a reference material is used, the requirements described in 4.2.1 also apply here.

5.5 Statistical analysis

5.5.1 Check of the within-laboratory standard deviation

Compute the average, \bar{y}_W , of the n test results and s_W , the estimate of the within-laboratory standard deviation σ_W , as follows:

$$\bar{y}_W = \frac{1}{n} \sum_{k=1}^n y_k \quad \dots (21)$$

$$s_W = \sqrt{\frac{1}{n-1} \sum_{k=1}^n (y_k - \bar{y}_W)^2} \quad \dots (22)$$

The test results shall be scrutinized for outliers using Grubbs' test as described in subclause 7.3.4 of ISO 5725-2:1994.

If the repeatability standard deviation, σ_r , of the standard measurement method is known, the estimate s_W can be assessed by the following procedure.

Compute the ratio

$$C'' = (s_W / \sigma_r)^2 \quad \dots (23)$$

and compare the value C'' with the critical value

$$C''_{\text{crit}} = \chi^2_{(1-\alpha)}(v) / v$$

where $\chi^2_{(1-\alpha)}(v)$ is the $(1-\alpha)$ -quantile of the χ^2 distribution with $v [= n - 1]$ degrees of freedom. Unless otherwise stated, α is assumed to be 0,05.

a) If $C'' \leq C''_{\text{crit}}$, s_W is not significantly larger than σ_r .

b) If $C'' > C''_{\text{crit}}$, s_W is significantly larger than σ_r .

In the former case, the repeatability standard deviation of the measurement method, σ_r , will be used for the assessment of the laboratory bias.

In the latter case, consideration should be given to repeating the experiment with verification at all steps that the standard measurement method is properly implemented.

5.5.2 Estimation of the laboratory bias

The estimate, $\hat{\Delta}$, of the laboratory bias Δ is given by

$$\hat{\Delta} = \bar{y}_W - \mu \quad \dots (24)$$

The variation of the estimate of the laboratory bias is due to the variation in the results of the measurement process and is expressed by its standard deviation computed as

$$\sigma_{\hat{\Delta}} = \sigma_r / \sqrt{n} \quad \dots (25)$$

in the case of a known repeatability standard deviation, or

$$s_{\hat{\Delta}} = s_W / \sqrt{n} \quad \dots (26)$$

in the case of an unknown repeatability standard deviation.

The 95 % confidence interval of the laboratory bias can be computed as

$$\hat{\Delta} - A_W \sigma_r \leq \Delta \leq \hat{\Delta} + A_W \sigma_r \quad \dots (27)$$

where A_W is as given in equation (20). If σ_r is unknown, its estimate s_r has to be used instead.

If this confidence interval covers the value zero, the laboratory bias is insignificant at the significance level $\alpha = 5 \%$; otherwise it is significant.

The laboratory bias is further considered in ISO 5725-6.

6 The report to, and the decisions to be taken by, the panel

6.1 Report by the statistical expert

Having completed the statistical analysis, the statistical expert shall write a report to be submitted to the panel. In this report the following information shall be given:

- a) a full account of the observations received from the operators and/or supervisors concerning the standard measurement method;
- b) a full account of the laboratories that have been rejected as outlying laboratories, together with the reasons for their rejection;
- c) a full account of any stragglers and/or outliers that have been identified, and whether these were explained and corrected, or discarded;
- d) a table of the final results of appropriate means and precision measures;
- e) a statement on whether the bias of the standard measurement method with respect to the accepted reference used is significant; if so, the estimated magnitude of the bias for each level shall be reported.

6.2 Decisions by the panel

The panel should then discuss the statistical expert's report and take decisions concerning the following questions.

- a) Are the discordant test results, if any, due to defects in the description of the measurement method?
- b) What action should be taken with respect to rejected outlying laboratories?
- c) Do the results of outlying laboratories and/or the comments received from the operators and supervisors indicate a need to improve the standard measurement method? If so, what are the improvements required?
- d) Do the results of the accuracy experiment justify the acceptability of the measurement method for adoption as a standard? What action is to be taken concerning its publication?

7 Utilization of trueness data

Refer to clause 7 of ISO 5725-1:1994.

Annex A (normative)

Symbols and abbreviations used in ISO 5725

a	Intercept in the relationship $s = a + bm$	k	Mandel's within-laboratory consistency test statistic
A	Factor used to calculate the uncertainty of an estimate	LCL	Lower control limit (either action limit or warning limit)
b	Slope in the relationship $s = a + bm$	m	General mean of the test property; level
B	Component in a test result representing the deviation of a laboratory from the general average (laboratory component of bias)	M	Number of factors considered in intermediate precision conditions
B_0	Component of B representing all factors that do not change in intermediate precision conditions	N	Number of iterations
$B_{(1)}, B_{(2)}, \text{ etc.}$	Components of B representing factors that vary in intermediate precision conditions	n	Number of test results obtained in one laboratory at one level (i.e. per cell)
c	Intercept in the relationship $\lg s = c + d \lg m$	p	Number of laboratories participating in the inter-laboratory experiment
C, C', C''	Test statistics	P	Probability
$C_{\text{crit}}, C'_{\text{crit}}, C''_{\text{crit}}$	Critical values for statistical tests	q	Number of levels of the test property in the interlaboratory experiment
CD_p	Critical difference for probability P	r	Repeatability limit
CR_p	Critical range for probability P	R	Reproducibility limit
d	Slope in the relationship $\lg s = c + d \lg m$	RM	Reference material
e	Component in a test result representing the random error occurring in every test result	s	Estimate of a standard deviation
f	Critical range factor	\hat{s}	Predicted standard deviation
$F_p(v_1, v_2)$	p -quantile of the F -distribution with v_1 and v_2 degrees of freedom	T	Total or sum of some expression
G	Grubbs' test statistic	t	Number of test objects or groups
h	Mandel's between-laboratory consistency test statistic	UCL	Upper control limit (either action limit or warning limit)
		W	Weighting factor used in calculating a weighted regression
		w	Range of a set of test results
		x	Datum used for Grubbs' test
		y	Test result

\bar{y}	Arithmetic mean of test results
$\bar{\bar{y}}$	Grand mean of test results
α	Significance level
β	Type II error probability
γ	Ratio of the reproducibility standard deviation to the repeatability standard deviation (σ_R/σ_r)
Δ	Laboratory bias
$\hat{\Delta}$	Estimate of Δ
δ	Bias of the measurement method
$\hat{\delta}$	Estimate of δ
λ	Detectable difference between two laboratory biases or the biases of two measurement methods
μ	True value or accepted reference value of a test property
ν	Number of degrees of freedom
ρ	Detectable ratio between the repeatability standard deviations of method B and method A
σ	True value of a standard deviation
τ	Component in a test result representing the variation due to time since last calibration
ϕ	Detectable ratio between the square roots of the between-laboratory mean squares of method B and method A
$\chi_p^2(\nu)$	p -quantile of the χ^2 -distribution with ν degrees of freedom

Symbols used as subscripts

C	Calibration-different
E	Equipment-different
i	Identifier for a particular laboratory
$I()$	Identifier for intermediate measures of precision; in brackets, identification of the type of intermediate situation
j	Identifier for a particular level (ISO 5725-2). Identifier for a group of tests or for a factor (ISO 5725-3)
k	Identifier for a particular test result in a laboratory i at level j
L	Between-laboratory (interlaboratory)
m	Identifier for detectable bias
M	Between-test-sample
O	Operator-different
P	Probability
r	Repeatability
R	Reproducibility
T	Time-different
W	Within-laboratory (intralaboratory)
1, 2, 3...	For test results, numbering in the order of obtaining them
(1), (2), (3)...	For test results, numbering in the order of increasing magnitude

Annex B (informative)

Example of an accuracy experiment

B.1 Description of the experiment

An accuracy experiment on the determination of manganese content in iron ores by an atomic absorption method was conducted by ISO/TC 102, *Iron ores*, using five test materials with the accepted reference values (μ) given in table B.1 (which were not disclosed to the laboratories). Each laboratory received two randomly selected bottles of test sample for each level and performed duplicate analyses on each bottle. The purpose of the two-bottle system was to confirm the absence of the between-bottle variation. The analysis was performed such that in the case where the absence of between-bottle variation is confirmed, the four analytical results can be considered as replicates under repeatability conditions. Analysis of the results showed that the between-bottle variation was indeed insignificant; the sample was considered to be homogeneous. Thus results from each laboratory can be considered as replicates under repeatability conditions. The analytical results are listed in table B.2. The laboratory means and variances for each of the five test materials are listed in table B.3.

B.2 Precision assessment

To assess the precision of the analytical method, the data were analysed by the procedure described in ISO 5725-2. The test results for each level are shown in figures B.1 to B.5.

The stragglers and outliers for both Cochran's and Grubbs' tests were identified and are listed in table B.4. The boxed points in figures B.1 to B.5 signify that the test results were identified as outliers. Table B.4 shows that seven laboratory results were identified as outliers; of these, five originated from two laboratories (Labs. 10 and 19). One laboratory result was identified as a straggler; it originated from the same laboratory (Lab. 10).

The h and k values are shown in figures B.6 and B.7. The h values (figure B.6) show clearly that laboratory 10 gets very low results; two of them (levels 2 and 3) were identified as outliers. It was therefore decided to discard the results from laboratory 10 completely; it should be the object of special attention, and the matter should be resolved. In addition, the data at level 1 of laboratory 7, identified as an outlier by Grubbs' test, were discarded. The k values (figure B.7) show that laboratories 10, 17 and 19 tend to get somewhat larger within-laboratory variation than the others. There again, appropriate action should be taken by investigating these laboratories, or, if necessary, by tightening the protocol of the measurement method. For the analysis, it was decided to discard the outliers identified by Cochran's test; i.e. the data at levels 3 and 5 of laboratory 19 and at level 5 of laboratory 17.

The repeatability and the reproducibility standard deviations were then computed excluding those data that were discarded. The results of this computation are summarized in table B.5 and plotted against the level in figure B.8. Figure B.8 shows that a linear function seems to be an appropriate relationship between the precisions and concentration levels. The linear regression equations of the repeatability and reproducibility standard deviations versus levels of concentration are:

$$s_r = 0,000\ 579 + 0,008\ 85m$$

$$s_R = 0,000\ 737 + 0,015\ 57m$$

B.3 Trueness assessment

The trueness of the measurement method was assessed by computing the 95 % confidence intervals of the bias of the measurement method using equation (19) and comparing them with zero (table B.5). Since at levels 3, 4 and 5 these confidence intervals cover the value zero, the bias of this measurement method is insignificant at the high con-

centration levels 3, 4 and 5 of manganese; since at levels 1 and 2 the confidence intervals do not cover zero, the bias is significant at the low concentration levels 1 and 2 of manganese.

B.4 Further analysis

Further information can be extracted from the data by carrying out supplementary analyses such as a regression analysis of \bar{y} versus μ .

Table B.1 — Manganese content in iron ores: Accepted reference values

Level	1	2	3	4	5
Accepted reference value μ (% Mn)	0,010 0	0,093 0	0,401 0	0,777 0	2,530 0

Table B.2 — Manganese content in iron ores: Analytical results as percentage Mn

Lab. No.	Bottle No.	Level									
		1		2		3		4		5	
1	1	0,011 8	0,012 1	0,088 0	0,087 5	0,408	0,407	0,791	0,791	2,584	2,560
	2	0,012 1	0,012 1	0,086 5	0,086 7	0,407	0,408	0,794	0,801	2,535	2,545
2	1	0,013 1	0,011 5	0,089 4	0,086 1	0,411	0,405	0,760	0,766	2,543	2,591
	2	0,011 5	0,011 5	0,088 7	0,086 7	0,406	0,399	0,766	0,783	2,516	2,567
3	1	0,011 8	0,011 2	0,086 4	0,084 9	0,410	0,403	0,752	0,767	2,526	2,463
	2	0,011 0	0,010 4	0,086 7	0,089 6	0,408	0,400	0,755	0,753	2,515	2,493
4	1	0,010 7	0,012 1	0,088 1	0,089 2	0,402	0,402	0,780	0,750	2,560	2,520
	2	0,011 4	0,012 1	0,086 1	0,087 4	0,404	0,402	0,777	0,750	2,600	2,520
5	1	0,012 0	0,012 8	0,090 4	0,090 4	0,404	0,400	0,775	0,775	2,470	2,510
	2	0,011 2	0,012 8	0,086 2	0,087 0	0,404	0,396	0,770	0,780	2,500	2,480
6	1	0,011 1	0,011 0	0,089 2	0,089 3	0,402	0,398	0,786	0,782	2,531	2,514
	2	0,011 0	0,011 1	0,090 0	0,086 4	0,408	0,404	0,780	0,772	2,524	2,494
7	1	0,008 8	0,009 5	0,089 3	0,089 5	0,390	0,390	0,754	0,762	2,510	2,521
	2	0,007 0	0,008 6	0,085 9	0,088 6	0,395	0,395	0,758	0,756	2,500	2,513
8	1	0,011 5	0,011 2	0,082 3	0,082 3	0,390	0,396	0,761	0,765	2,501	2,499
	2	0,011 3	0,011 3	0,082 8	0,082 9	0,400	0,389	0,770	0,766	2,507	2,490
9	1	0,012 3	0,012 0	0,086 2	0,086 6	0,414	0,414	0,765	0,765	2,523	2,520
	2	0,011 7	0,011 8	0,086 5	0,087 6	0,411	0,414	0,765	0,765	2,521	2,508
10	1	0,009 5	0,008 6	0,078 0	0,072 0	0,390	0,370	0,746	0,730	2,530	2,580
	2	0,009 2	0,008 4	0,078 0	0,073 0	0,392	0,374	0,750	0,738	2,510	2,610
11	1	0,012 5	0,012 5	0,090 0	0,089 0	0,405	0,395	0,790	0,780	2,520	2,520
	2	0,013 0	0,012 5	0,089 0	0,089 5	0,400	0,405	0,785	0,790	2,530	2,520
12	1	0,012 5	0,013 0	0,088 5	0,089 0	0,405	0,395	0,790	0,780	2,535	2,525
	2	0,011 5	0,013 0	0,089 0	0,087 5	0,405	0,390	0,775	0,790	2,550	2,495
13	1	0,012 5	0,011 6	0,084 2	0,083 2	0,399	0,399	0,784	0,777	2,523	2,523
	2	0,012 1	0,011 6	0,083 2	0,082 8	0,398	0,399	0,782	0,777	2,527	2,537
14	1	0,011 6	0,012 0	0,089 8	0,089 0	0,418	0,416	0,797	0,800	2,602	2,602
	2	0,009 8	0,011 6	0,090 0	0,090 2	0,415	0,415	0,801	0,790	2,592	2,602
15	1	0,010 8	0,011 2	0,087 1	0,086 0	0,399	0,400	0,775	0,774	2,488	2,495
	2	0,011 2	0,011 1	0,088 3	0,086 1	0,397	0,401	0,783	0,773	2,503	2,485
16	1	0,010 9	0,010 8	0,084 6	0,085 8	0,392	0,400	0,779	0,769	2,528	2,516
	2	0,011 1	0,011 0	0,084 9	0,085 5	0,396	0,397	0,751	0,753	2,528	2,525
17	1	0,010 0	0,011 0	0,084 9	0,088 0	0,409	0,410	0,766	0,794	2,571	2,380
	2	0,010 0	0,010 0	0,083 0	0,089 0	0,392	0,402	0,755	0,775	2,429	2,488
18	1	0,011 7	0,010 2	0,088 0	0,088 1	0,405	0,404	0,771	0,773	2,520	2,511
	2	0,012 5	0,010 3	0,086 8	0,088 2	0,402	0,403	0,778	0,763	2,514	2,503
19	1	0,009 9	0,012 8	0,094 5	0,090 5	0,398	0,375	0,770	0,767	2,483	2,351
	2	0,011 8	0,012 8	0,092 4	0,088 4	0,418	0,382	0,799	0,760	2,485	2,382

Table B.3 — Manganese content in iron ores: Laboratory means and laboratory variances

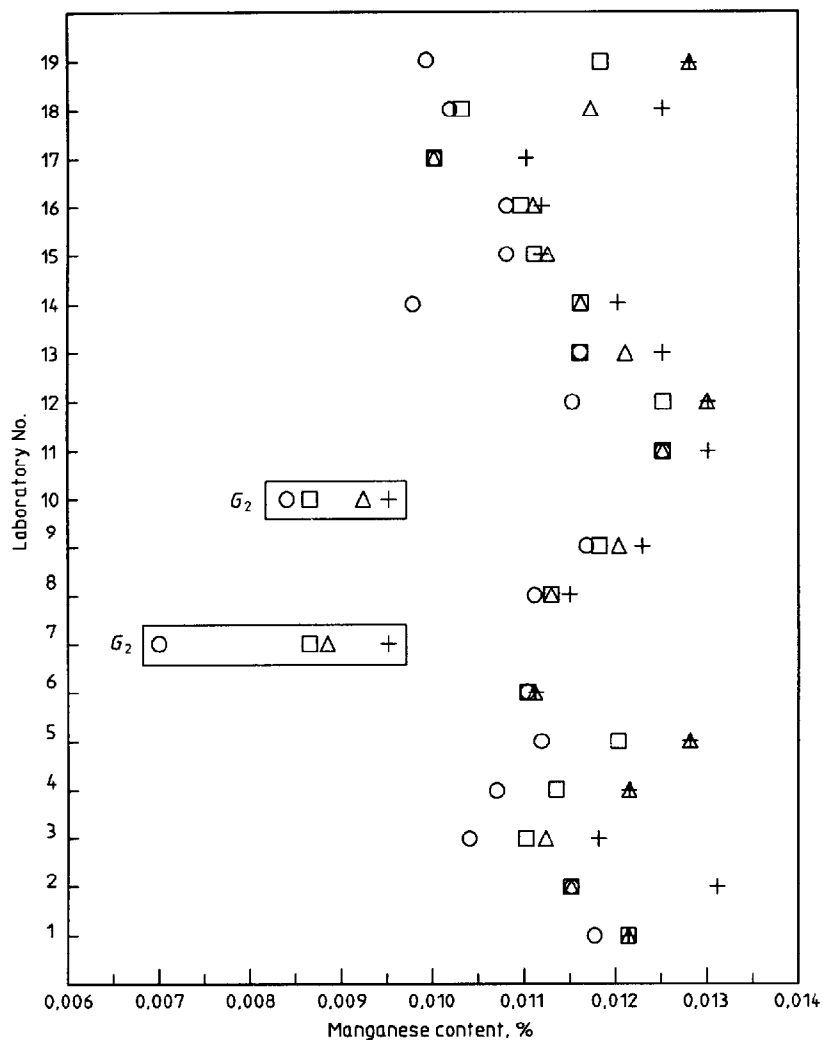
Lab. No.	Level				
	1	2	3	4	5
Laboratory mean					
1	0,012 03	0,087 18	0,407 50	0,794 25	2,556 00
2	0,011 90	0,087 73	0,405 25	0,768 75	2,554 25
3	0,011 10	0,086 90	0,405 25	0,756 75	2,499 25
4	0,011 58	0,087 70	0,402 50	0,764 25	2,550 00
5	0,012 20	0,088 50	0,401 00	0,775 00	2,490 00
6	0,011 05	0,088 73	0,403 00	0,780 00	2,515 75
7	0,008 48	0,088 33	0,392 50	0,757 50	2,511 00
8	0,011 33	0,082 58	0,393 75	0,765 50	2,499 25
9	0,011 95	0,086 73	0,413 25	0,765 00	2,518 00
10	0,008 93	0,075 25	0,381 50	0,741 00	2,557 50
11	0,012 63	0,089 38	0,401 25	0,786 25	2,522 50
12	0,012 50	0,088 50	0,398 75	0,783 75	2,526 25
13	0,011 95	0,083 35	0,398 75	0,780 00	2,527 50
14	0,011 25	0,089 75	0,416 00	0,797 00	2,599 50
15	0,011 08	0,086 88	0,399 25	0,776 25	2,492 75
16	0,010 95	0,085 20	0,396 25	0,763 00	2,524 25
17	0,010 25	0,086 23	0,403 25	0,772 50	2,467 00
18	0,011 18	0,087 78	0,403 50	0,771 25	2,512 00
19	0,011 83	0,091 45	0,393 25	0,774 00	2,425 25
Laboratory variance					
1	$0,225\ 0 \times 10^{-7}$	$0,489\ 2 \times 10^{-6}$	$0,333\ 3 \times 10^{-6}$	$0,222\ 5 \times 10^{-4}$	$0,454\ 0 \times 10^{-3}$
2	$0,640\ 0 \times 10^{-6}$	$0,248\ 2 \times 10^{-5}$	$0,242\ 5 \times 10^{-4}$	$0,982\ 5 \times 10^{-4}$	$0,103\ 4 \times 10^{-2}$
3	$0,333\ 3 \times 10^{-6}$	$0,386\ 0 \times 10^{-5}$	$0,209\ 2 \times 10^{-4}$	$0,482\ 5 \times 10^{-4}$	$0,772\ 2 \times 10^{-3}$
4	$0,449\ 2 \times 10^{-6}$	$0,168\ 7 \times 10^{-5}$	$0,100\ 0 \times 10^{-5}$	$0,272\ 2 \times 10^{-3}$	$0,146\ 7 \times 10^{-2}$
5	$0,586\ 7 \times 10^{-6}$	$0,492\ 0 \times 10^{-5}$	$0,146\ 7 \times 10^{-4}$	$0,166\ 7 \times 10^{-4}$	$0,333\ 3 \times 10^{-3}$
6	$0,333\ 3 \times 10^{-8}$	$0,252\ 9 \times 10^{-5}$	$0,173\ 3 \times 10^{-4}$	$0,346\ 7 \times 10^{-4}$	$0,258\ 9 \times 10^{-3}$
7	$0,111\ 6 \times 10^{-5}$	$0,276\ 3 \times 10^{-5}$	$0,833\ 3 \times 10^{-5}$	$0,116\ 7 \times 10^{-4}$	$0,753\ 3 \times 10^{-4}$
8	$0,158\ 3 \times 10^{-7}$	$0,102\ 5 \times 10^{-6}$	$0,269\ 2 \times 10^{-4}$	$0,136\ 7 \times 10^{-4}$	$0,495\ 8 \times 10^{-4}$
9	$0,700\ 0 \times 10^{-7}$	$0,369\ 2 \times 10^{-6}$	$0,225\ 0 \times 10^{-5}$	0	$0,460\ 0 \times 10^{-4}$
10	$0,262\ 5 \times 10^{-6}$	$0,102\ 5 \times 10^{-4}$	$0,123\ 7 \times 10^{-3}$	$0,786\ 7 \times 10^{-4}$	$0,209\ 2 \times 10^{-2}$
11	$0,625\ 0 \times 10^{-7}$	$0,229\ 2 \times 10^{-6}$	$0,229\ 2 \times 10^{-4}$	$0,229\ 2 \times 10^{-4}$	$0,250\ 0 \times 10^{-4}$
12	$0,500\ 0 \times 10^{-6}$	$0,500\ 0 \times 10^{-6}$	$0,562\ 5 \times 10^{-4}$	$0,562\ 5 \times 10^{-4}$	$0,539\ 6 \times 10^{-3}$
13	$0,190\ 0 \times 10^{-6}$	$0,356\ 7 \times 10^{-6}$	$0,250\ 0 \times 10^{-6}$	$0,126\ 7 \times 10^{-4}$	$0,436\ 7 \times 10^{-4}$
14	$0,970\ 0 \times 10^{-6}$	$0,276\ 7 \times 10^{-6}$	$0,200\ 0 \times 10^{-5}$	$0,246\ 7 \times 10^{-4}$	$0,250\ 0 \times 10^{-4}$
15	$0,358\ 3 \times 10^{-7}$	$0,114\ 9 \times 10^{-5}$	$0,291\ 7 \times 10^{-5}$	$0,209\ 2 \times 10^{-4}$	$0,642\ 5 \times 10^{-4}$
16	$0,166\ 7 \times 10^{-7}$	$0,300\ 0 \times 10^{-6}$	$0,109\ 2 \times 10^{-4}$	$0,178\ 7 \times 10^{-3}$	$0,322\ 5 \times 10^{-4}$
17	$0,250\ 0 \times 10^{-6}$	$0,766\ 9 \times 10^{-5}$	$0,689\ 2 \times 10^{-4}$	$0,272\ 3 \times 10^{-3}$	$0,675\ 7 \times 10^{-2}$
18	$0,124\ 9 \times 10^{-5}$	$0,429\ 2 \times 10^{-6}$	$0,166\ 7 \times 10^{-5}$	$0,389\ 2 \times 10^{-4}$	$0,500\ 0 \times 10^{-4}$
19	$0,186\ 9 \times 10^{-5}$	$0,680\ 3 \times 10^{-5}$	$0,364\ 9 \times 10^{-3}$	$0,295\ 3 \times 10^{-3}$	$0,476\ 3 \times 10^{-2}$

Table B.4 — Manganese content in iron ores: Outliers and stragglers

Level	Lab.	Calculated statistic ¹⁾	Critical value ¹⁾
List of outliers ($\alpha = 0,01$)			
1	7	$G2 = 0,295$	$G2(19) = 0,339\ 8$
2	10	$G1 = 3,305$	$G1(19) = 2,968$
3	19	$C = 0,474$	$C(4,19) = 0,276$
4	10	$C = 0,305$	$C(4,18) = 0,288$
5	—	—	—
5	17	$C = 0,358$	$C(4,19) = 0,276$
	19	$C = 0,393$	$C(4,18) = 0,288$
List of stragglers ($\alpha = 0,05$)			
1	—	—	—
2	—	—	—
3	—	—	—
4	—	—	—
5	10	$C = 0,284$	$C(4,17) = 0,250$
1) C = Cochran's test $G1$ = Grubbs' test for one outlying observation $G2$ = Grubbs' test for two outlying observations			

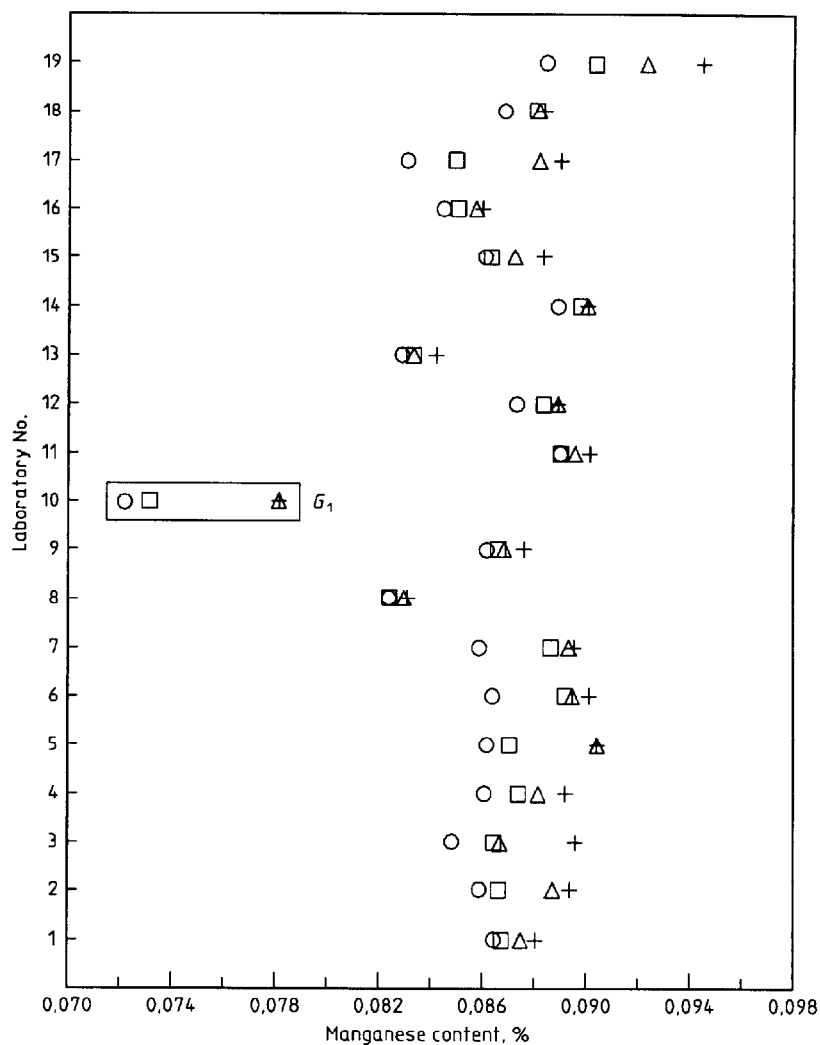
Table B.5 — Manganese content in iron ores: Estimation of repeatability and reproducibility standard deviations and bias of the measurement method

	Level				
	1	2	3	4	5
n	4	4	4	4	4
p	17	18	17	18	16
s_r	0,000 65	0,001 43	0,004 07	0,008 95	0,018 15
s_R	0,000 84	0,002 48	0,007 06	0,013 85	0,032 46
γ	1,29	1,73	1,73	1,54	1,79
A	0,352 8	0,399 9	0,411 7	0,383 0	0,428 7
As_R	0,000 296	0,000 991	0,002 906	0,005 301	0,013 916
\bar{y}	0,011 6	0,087 4	0,402 4	0,773 9	2,524 9
μ	0,010 0	0,093 0	0,401 0	0,777 0	2,530 0
$\hat{\delta}$	0,001 6	− 0,005 6	0,001 4	− 0,003 1	− 0,005 1
$\hat{\delta} - As_R$	0,001 3	− 0,006 6	− 0,001 5	− 0,008 4	− 0,019 0
$\hat{\delta} + As_R$	0,001 9	− 0,004 6	0,004 3	0,002 2	0,008 8



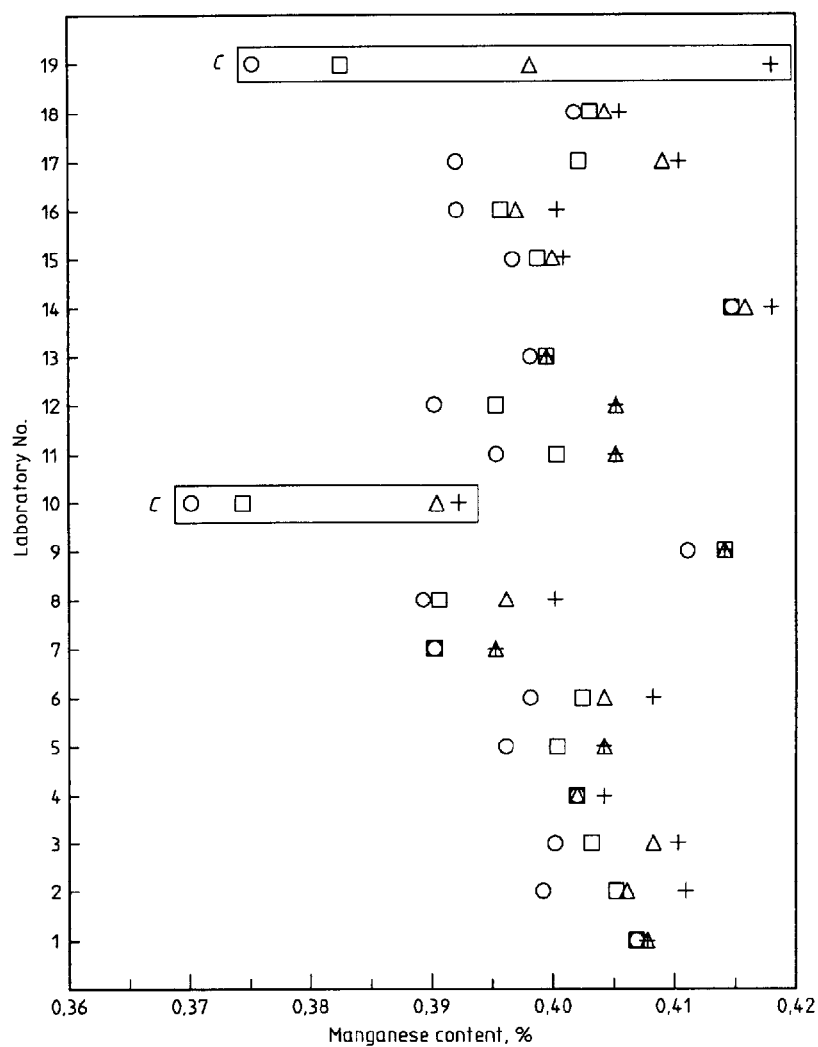
NOTE — Boxed points signify that the test results were identified as outliers by Grubbs' test for two outlying observations (G_2).

Figure B.1 — Manganese content in iron ores: Test results at level 1



NOTE — Boxed points signify that the test results were identified as outliers by Grubbs' test for one outlying observation (G_1).

Figure B.2 — Manganese content in iron ores: Test results at level 2



NOTE — Boxed points signify that the test results were identified as outliers by Cochran's test (C).

Figure B.3 — Manganese content in iron ores: Test results at level 3

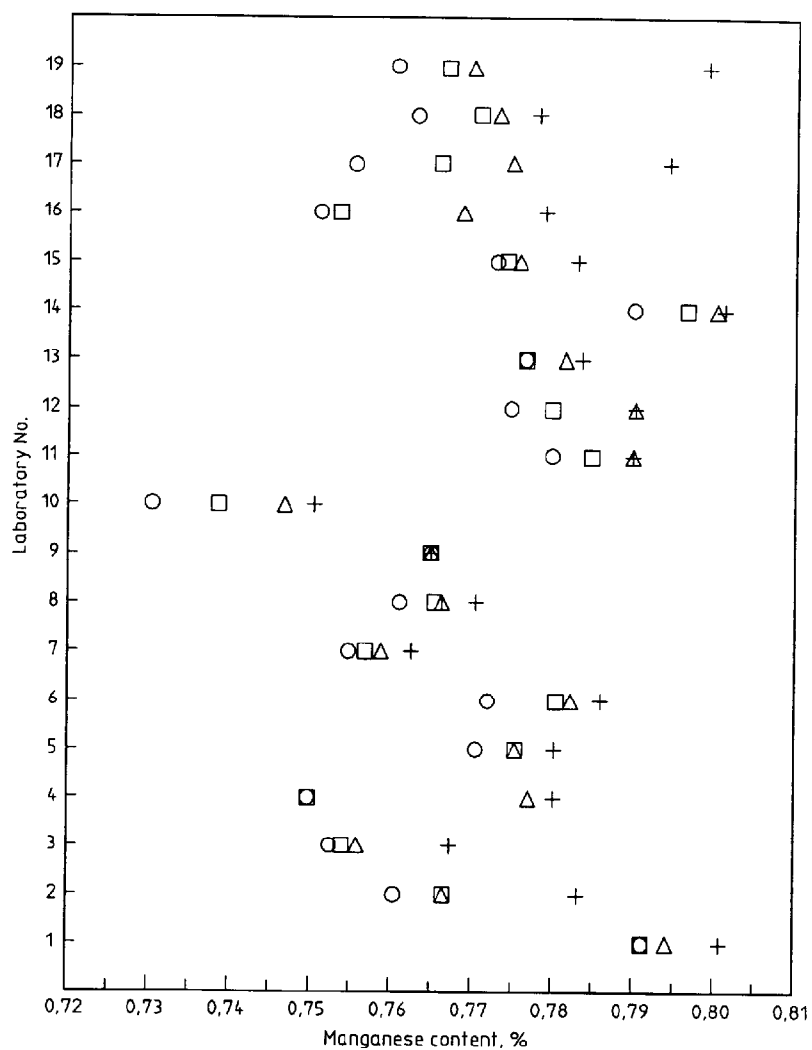
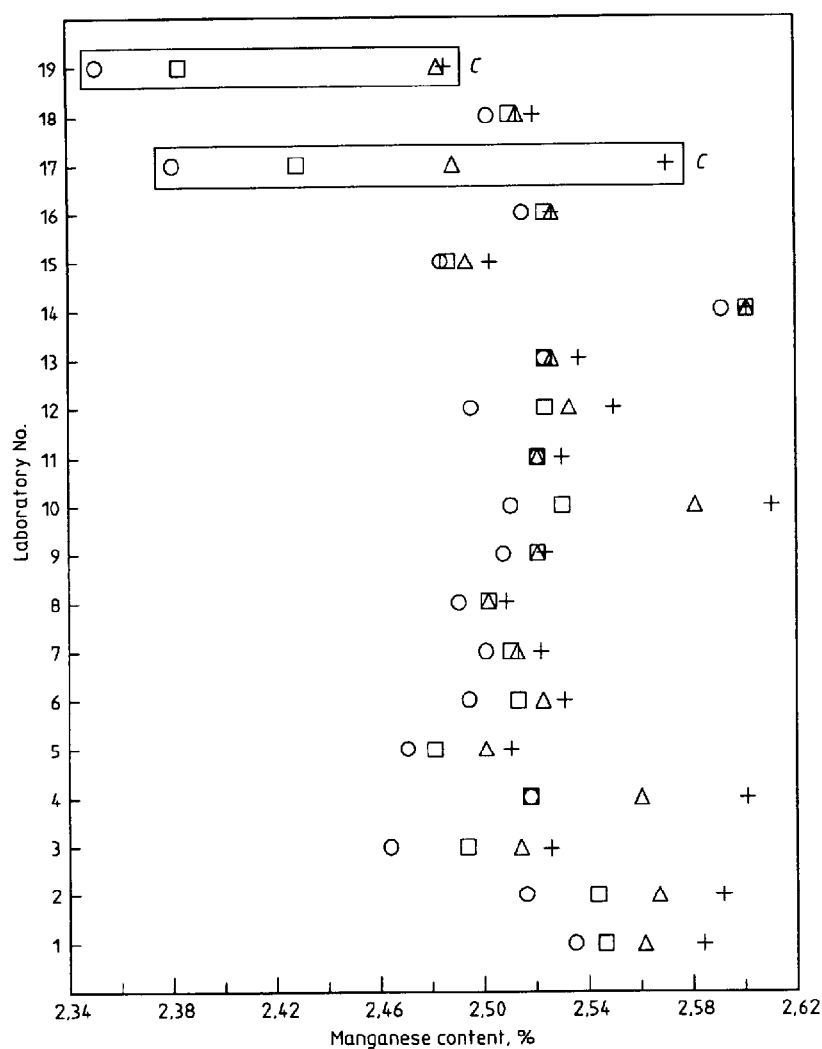


Figure B.4 — Manganese content in iron ores: Test results at level 4



NOTE — Boxed points signify that the test results were identified as outliers by Cochran's test (C).

Figure B.5 — Manganese content in iron ores: Test results at level 5

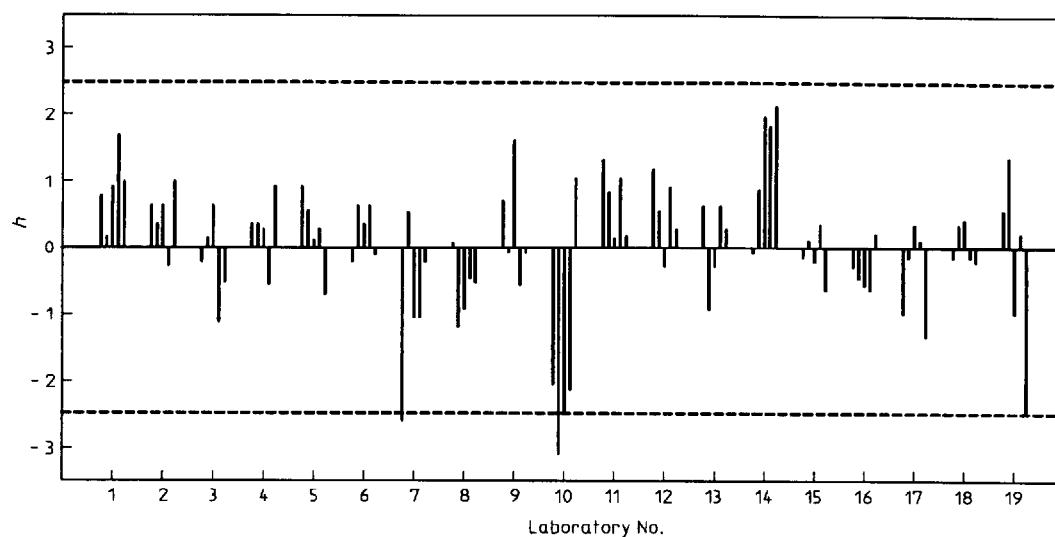


Figure B.6 — Manganese content in iron ores: h values grouped by laboratories

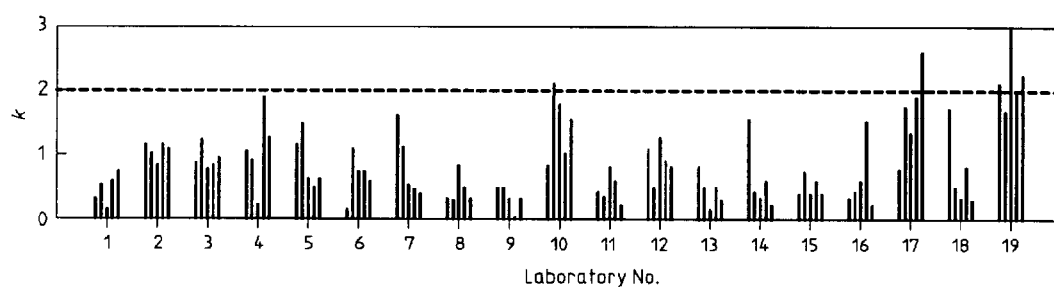


Figure B.7 — Manganese content in iron ores: k values grouped by laboratories

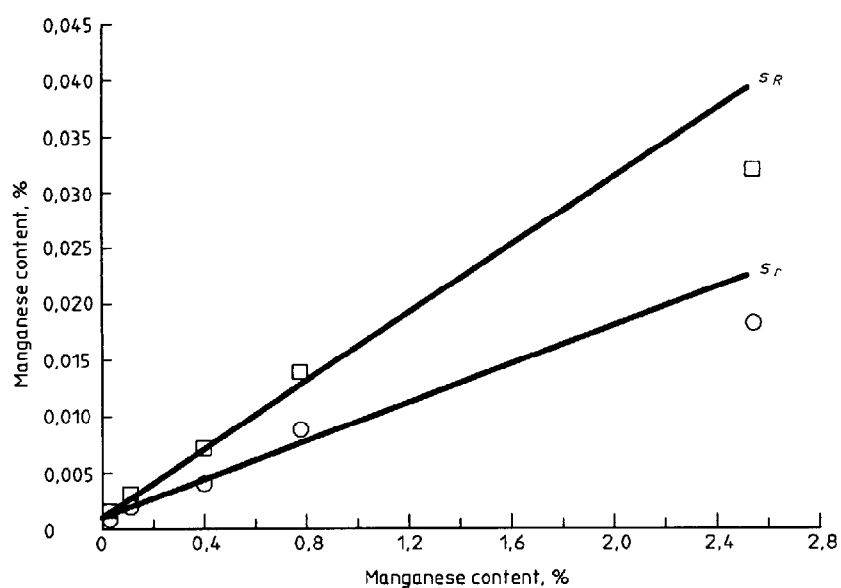


Figure B.8 — Manganese content in iron ores: Repeatability and reproducibility standard deviations as linear functions of the concentration level m

Annex C

(informative)

Derivation of equations

C.1 Equations (5) and (6) (see 4.5)

The minimum number of laboratories, p , and of test results, n , are calculated to satisfy the two following conditions:

- a) the test should be able to detect that the bias is equal to zero with the probability $1 - \alpha = 0,95$;
- b) the test should be able to detect a predetermined magnitude of bias, δ_m , with the probability $1 - \beta = 0,95$.

The first condition is actually developed in 4.7.2, where the confidence interval for the bias of the measurement method, δ , is used to carry out a statistical test of the null hypothesis that the bias is equal to zero ($H_0: \delta = 0$) against the alternative hypothesis that the bias is unequal to zero ($H_1: \delta \neq 0$).

An equivalent form of this test would be to compare the absolute value of the estimate of the bias of the measurement method

$$|\hat{\delta}| = |\bar{y} - \mu|$$

with a critical value K , and reject H_0 ($\delta = 0$) if $|\hat{\delta}| > K$ [and not reject H_0 ($\delta = 0$) if $|\hat{\delta}| \leq K$].

K can be computed using the requirement that the probability of rejecting H_0 , if it is true, shall be equal to the chosen significance level $\alpha = 5\%$:

$$P(|\hat{\delta}| > K | \delta = 0) = \alpha = 0,05$$

$$P(|\hat{\delta}| \leq K | \delta = 0) = 1 - \alpha = 0,95$$

$$= \Phi\left(\frac{K}{\sqrt{V(\hat{\delta})}}\right) - \Phi\left(-\frac{K}{\sqrt{V(\hat{\delta})}}\right)$$

$$= 2\Phi\left(\frac{K}{\sqrt{V(\hat{\delta})}}\right) - 1$$

$$\Phi\left(\frac{K}{\sqrt{V(\hat{\delta})}}\right) = 0,975$$

$$\frac{K}{\sqrt{V(\hat{\delta})}} = u_{0,975} = 1,960$$

$$K = 1,960\sqrt{V(\hat{\delta})} \quad \dots (C.1)$$

where

$\Phi(\cdot)$ is the cumulative distribution function of the standard normal distribution;

u_p is the p -quantile of the standard normal distribution;

$V(\hat{\delta})$ is the variance of the estimate of the bias of the measurement method.

$$\begin{aligned} V(\hat{\delta}) &= V(\bar{y} - \mu) = V(\bar{y}) \\ &= \frac{\sigma_L^2}{p} + \frac{\sigma_r^2}{pn} \\ &= \frac{\sigma_R^2 - \sigma_r^2}{p} + \frac{\sigma_r^2}{pn} \\ &= \frac{n(\sigma_R^2 - \sigma_r^2/\gamma^2) + \sigma_r^2/\gamma^2}{pn} \\ &= \left(\frac{n(\gamma^2 - 1) + 1}{\gamma^2 pn} \right) \sigma_R^2 \end{aligned}$$

where σ_L^2 is the between-laboratory variance so that

$$\sigma_R^2 = \sigma_L^2 + \sigma_r^2 \text{ and}$$

$$\gamma = \sigma_R/\sigma_r$$

The second condition is that the test should be able to detect the predetermined magnitude of bias, δ_m , with the probability $1 - \beta = 0,95$:

$$P(|\hat{\delta}| > K | \delta = \delta_m) = 1 - \beta = 0,95$$

$$P(|\hat{\delta}| \leq K|\delta = \delta_m) = \beta = 0,05$$

$$= P\left(\frac{\hat{\delta} - \delta_m}{\sqrt{V(\hat{\delta})}} \leq \frac{K - \delta_m}{\sqrt{V(\hat{\delta})}}\right) = \Phi\left(\frac{K - \delta_m}{\sqrt{V(\hat{\delta})}}\right)$$

$$\frac{K - \delta_m}{\sqrt{V(\hat{\delta})}} = u_{0,05} = -1,645$$

$$K = \delta_m - 1,645\sqrt{V(\hat{\delta})} \quad \dots (C.2)$$

Equating the two expressions (C.1 and C.2) for K gives

$$1,960\sqrt{V(\hat{\delta})} = \delta_m - 1,645\sqrt{V(\hat{\delta})}$$

$$(1,960 + 1,645)\sqrt{V(\hat{\delta})} = \delta_m$$

$$\left(1 + \frac{1,645}{1,960}\right)1,960\sqrt{V(\hat{\delta})} = \delta_m$$

$$\left(1 + \frac{1,645}{1,960}\right)A\sigma_R = \delta_m$$

$$A\sigma_R = \frac{\delta_m}{1,84}$$

C.2 Equations (19) and (20) (see 5.3)

These equations follow immediately if in the preceding derivation (C.1) δ , δ_m , $\hat{\delta}$, $V(\hat{\delta})$ and A are replaced by Δ , Δ_m , $\hat{\Delta}$, $V(\hat{\Delta})$ and A_W , respectively, and the expression for $V(\hat{\delta})$ is replaced by the expression

$$V(\hat{\Delta}) = \sigma_r^2/n$$

Annex D

(informative)

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